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#### Choquet integral in decision making and metric learning

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#### **Basics and objectives:**

- Using Choquet integral in two types of applications decision and metric learning (reidentification)
- Distances
- and distribution
  - (for non-additive measures)

#### 1. Preliminaries

- Choquet integral: mathematical perspective
  - Non-additive measures
  - $\circ$  Now we need an integral
- Choquet integral: Application perspective
  - $\circ$  Aggregation operators and CI in decision: MCDM
  - $\circ$  Aggregation operators and CI in reidentification: risk assessment
  - $\circ$  Zooming out
- 2. Distances in classification (filling the gaps)
- 3. Distributions

# **Choquet integral:** a mathematical introduction

# **Non-additive measures**

### **Definitions: measures**

#### Additive measures.

(X, A) a measurable space; then, a set function μ is an additive measure if it satisfies
(i) μ(A) ≥ 0 for all A ∈ A,
(ii) μ(X) ≤ ∞
(iii) for every countable sequence A<sub>i</sub> (i ≥ 1) of A that is pairwise disjoint (i.e,. A<sub>i</sub> ∩ A<sub>j</sub> = Ø when i ≠ j)

$$\mu(\bigcup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} \mu(A_i)$$

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$$\mu(\bigcup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} \mu(A_i)$$

Finite case:  $\mu(A \cup B) = \mu(A) + \mu(B)$  for disjoint A, B

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- Lebesgue measure. Unique measure  $\lambda$  s.t.  $\lambda([a,b])=b-a$  for every finite interval [a,b]
- Probability. When  $\mu(X) = 1$ .
- Or just price ...



#### Non-additive measures

- $(X, \mathcal{A})$  a measurable space, a non-additive (fuzzy) measure  $\mu$  on  $(X, \mathcal{A})$  is a set function  $\mu : \mathcal{A} \to [0, 1]$  satisfying the following axioms:
- (i)  $\mu(\emptyset) = 0$ ,  $\mu(X) = 1$  (boundary conditions) (ii)  $A \subseteq B$  implies  $\mu(A) \le \mu(B)$  (monotonicity)

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- Naturally, additivity implies monotonicity

• E.g.,  $B = A \cup C$  (with  $A \cap C = \emptyset$ ) then  $\mu(B) = \mu(A) + \mu(C) \ge \mu(A)$ 

• But in non-additive measures, we allow

$$\mu(B = A \cup C) < \mu(A) + \mu(C)$$
$$\mu(B = A \cup C) > \mu(A) + \mu(C)$$

As e.g.,  $\mu(B)=0.5<\mu(A)+\mu(C)=0.3+0.4=0.7$  A way to represent interactions

- Non-additive measures. Price
  - $\circ$  When we have a discount, for disjoints A and B, we have

 $\mu(A \cup B) < \mu(A) + \mu(B)$  but  $\mu(A \cup B) \ge \mu(A)$ 



• There quite a large number of families of measures

- Non-additive measures. Distorted probabilities
  - $\circ m : \mathbb{R}^+ \to \mathbb{R}^+$  a continuous and increasing function such that m(0) = 0; P be a probability.

$$\mu_{m,P}(A) = m(P(A)) \tag{1}$$



Used in economics: Prospect theory (Kahneman and Tversky, 1979).
 Small probabilities tend to be overestimated, while large ones, underestimated.

- Non-additive measures. Distorted Lebesgue
  - $\circ m : \mathbb{R}^+ \to \mathbb{R}^+$  a continuous and increasing function such that m(0) = 0;  $\lambda$  be the Lebesgue measure.

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• If  $m(x) = x^2$ , then  $\mu_m(A) = (\lambda(A))^2$ • If  $m(x) = x^p$ , then  $\mu_m(A) = (\lambda(A))^p$ 



- Non-additive measures. A large number of families
  - $\circ$  Sugeno λ-measures:  $\mu(A \cup B) = \mu(A) + \mu(B) + \lambda \mu(A) \mu(B)$  (λ > −1)
  - $\circ$  For  $\mathcal P$  a non empty set of probability measures, the upper and lower probabilities
    - $\triangleright \bar{P}(A) = \sup_{P \in \mathcal{P}} P(A)$
    - $\triangleright \underline{P}(A) = \inf_{P \in \mathcal{P}} P(A)$ 
      - (dual in the sense:  $\overline{P}(A) = 1 \underline{P}(A^c)$ )
- m-dimensional distorted probabilities (NT/NT, 2005, 2011, 2012, 2018)



Vicenç Torra; Choquet integral in decision making and metric learning

# Now we need an integral

• Additive measure: the way you add areas does not change<sup>1</sup> results



- Riemann integral (a) vs Lebesgue integral (c)
  - Riemann sum:  $\sum_{I \in \mathcal{C}} f(x(I)) * \mu(I)$ ( $\mathcal{C}$  non-overlapping collection, x(I) an element of I)
  - Lebesgue sum:  $\sum_{a_i \in Range(f)} (a_i a_{i-1}) \mu(\Gamma(a_i))$ where  $\Gamma(a) := \{x | f(x) \ge a\}$

<sup>&</sup>lt;sup>1</sup>Well, if it is calculable

• Lebesgue integral

$$\int f d\mu := \int_0^\infty \mu_f(r) dr$$
 where  $\mu_f(r) = \mu(\{x | f(x) \ge r\})$ 

- Choquet integral (Choquet, 1954):
  - $\circ \mu$  a non-additive measure, f a measurable function. The Choquet integral of f w.r.t.  $\mu$ , where  $\mu_f(r) := \mu(\{x | f(x) > r\})$ :

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$$(C)\int fd\mu := \int_0^\infty \mu_f(r)dr.$$

- Properties.
  - When the measure is additive, this is the Lebesgue integral (standard integral)

#### Choquet integral. Discrete version

•  $\mu$  a non-additive measure, f a measurable function. The Choquet integral of f w.r.t.  $\mu,$ 

$$(C)\int fd\mu = \sum_{i=1}^{N} [f(x_{s(i)}) - f(x_{s(i-1)})]\mu(A_{s(i)}),$$

where  $f(x_{s(i)})$  indicates that the indices have been permuted so that  $0 \leq f(x_{s(1)}) \leq \cdots \leq f(x_{s(N)}) \leq 1$ , and where  $f(x_{s(0)}) = 0$  and  $A_{s(i)} = \{x_{s(i)}, \dots, x_{s(N)}\}.$ 

#### • Choquet integral. Example:

• Distorted probability  $\mu_m(A) = m(P(A))$  (with m(0) = 0, m(1) = 1)  $CI_{\mu_m}(f)$ : (a)  $\rightarrow$  max, (b)  $\rightarrow$  median, (c)  $\rightarrow$  min, (d)  $\rightarrow$  mean (expectation)



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Opper and lower probabilities: bounds for expectations
 CI<sub>P</sub>(f) ≤ inf<sub>P</sub> E<sub>P</sub>(f) ≤ sup<sub>P</sub> E<sub>P</sub>(f) ≤ CI<sub>P̄</sub>(f)
 (C) ∫ χ<sub>A</sub>dμ = μ(A)

# Application I Aggregation operators & Choquet integral in Decision

# MCDM: Aggregation for (numerical) utility functions

• Decision, utility functions

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- Decision making process:
- Modelling=Criteria + Utilities, aggregation, selection

	Number of	Security	Price	Confort	trunk
	seats				
Ford T	0	20	0	20	0
Seat 600	60	0	100	0	50
Simca 1000	100	30	100	50	70
VW Beetle	80	50	30	70	100
Citroën Acadiane	20	40	60	40	0

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 Decision making process:

Modelling, aggregation =  $\mathbb{C}$ , selection

	Seats	Security	Price	Comfort	trunk	$\mathbb{C} = AM$
Ford T	0	20	0	20	0	8
Seat 600	60	0	100	0	50	42
Simca 1000	100	30	100	50	70	70
VW	80	50	30	70	100	66
Citr. Acadiane	20	40	60	40	0	32

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when  $a_i \leq b_i$  it is clear that  $a \leq b$ 

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Aggregation operators are appropriate because they satisfy monotonicity

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Modelling, aggregation, selection=order,first

- The function of aggregation functions
  - Different aggregations lead to different orders (in the PF)
  - Aggregation establishes which points are equivalent

Different aggregations, lead to different curves of points (level curves)



- Aggregation functions and different level curves
  - Arithmetic mean
  - Geometric mean, Harmonic mean, ...
  - Weighted mean
  - OWA, ...

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  - Choquet integral (generalization of the AM, WM, OWA)
    - ▷ to represent interactions between criteria
    - > non-independent criteria allowed

- Aggregation functions and parameters
  - Arithmetic mean: no parameters
  - Geometric mean, Harmonic mean, ...: : no parameters
  - Weighted mean: weighting vector
  - OWA, ...: weighting vector
  - Choquet integral (generalization of the AM, WM, OWA) a measure
    - ▷ to represent interactions between criteria

*w*(*security*,*price*,*confort*) > (or <) *w*(*security*)+*w*(*price*)+*w*(*confort*)

> non-independent criteria allowed

 $\mu(\{c_1, c_2\}) \neq \mu(\{c_1\}) + \mu(\{c_2\})$ 

 $\triangleright (C) \int \chi_A d\mu = \mu(A)$ 

# MCDM: What fuzzy measures (and CI) can represent?

- Choquet integral can, and WM/Probability model cannot
  - An element/criteria is added into the set, and the preference is reversed

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  - $\circ$  An element/criteria is added into the set, and

the preference is reversed

• Example. Buying a house.

When public transport is available, the preference changes<sup>2</sup>

- If there is no bus I prefer a public library than a restaurant, but if there is a bus then I instead prefer the restaurant near.
- $\triangleright \text{ Mathematically, with } B=Bus, \ R=Restaurant, \ L=Library \text{ we have } \\ \mu(\{R\}) \leq \mu(\{L\}) \text{ but } \mu(\{R,B\}) \geq \mu(\{L,B\})$

<sup>&</sup>lt;sup>2</sup>Ellesberg's paradox.

## MCDM: Learn/identify the parameters (e.g. the measures)

#### • Available information?

0	Find measure	es from Seats	outcome Security	e: <mark>col</mark> Price	<mark>umn vect</mark> Comfort	or with trunk	$\begin{array}{c} \text{outcome} \\ \mathbb{C} = CI_{\mu} \end{array}$
	Seat 600	60	0	100	0	50	42
	Simca 1000	100	30	100	50	70	70

0	$\sim$ Find measures from preferences – (partial) order <: $S = \{(r_i, t_i)\}_i$						
		Seats	Security	Price	Comfort	trunk	$\mathbb{C} = CI_{\mu}$
	Seat 600	60	0	100	0	50	4th
	Simca 1000	100	30	100	50	70	1st

- Available information?
  - Measures from outcome: a column vector  $\Rightarrow \min \sum (\mathbb{C}_P(a_r) o_r)^2$
  - Measures from preferences (partial) order <:  $S = \{(r_i, t_i)\}_i$ 
    - $\triangleright \text{ Formulation: Find } \mu \text{ such that, for all } (r,t) \in S, \text{ it follows that} \\ \mathbb{C}_P(\text{evaluation-car } r) > \mathbb{C}_P(\text{evaluation-car } t) \\ \text{or, with } a_r \text{ and } a_s \text{ for rows } r \text{ and } s, \\ \mathbb{C}_P(a_{r1}, \dots, a_{rn}) > \mathbb{C}_P(a_{t1}, \dots, a_{tn}) \\ \text{Unfortunately, often, no solution: minimize failures } y_{(r,t)} \geq 0 \\ \mathbb{C}_P(a_{r1}, \dots, a_{rn}) \mathbb{C}_P(r_{t1}, \dots, a_{tn}) + y_{(r,t)} > 0. \end{cases}$

- Available information?
  - Measures from outcome: a column vector  $\Rightarrow \min \sum (\mathbb{C}_P(a_r) o_r)^2$
  - Measures from preferences (partial) order <:  $S = \{(r_i, t_i)\}_i$ 
    - $\triangleright$  Formulation: Find  $\mu$  such that, for all  $(r,t)\in S$  , it follows that

Minimize  $\sum_{(r,t)\in S} y_{(r,t)}$ Subject to  $\mathbb{C}_P(a_{r1},\ldots,a_{rn}) - \mathbb{C}_P(a_{t1},\ldots,a_{tn}) + \begin{array}{l} y_{(r,t)} > 0 \\ y_{(r,t)} \ge 0 \end{array}$ logical constraints on P

- Aggregation and selection
  - $\circ\,$  Selection of the one with maximum value of  $\mathbb{C}=CI$  with  $\mu$  (maximum distance to nadir worst combination)

$$d((a_1,\ldots,a_n),(0,\ldots,0))$$

• Selection of the one with minimum distance to ideal  $d((a_1, \ldots, a_n), (100, \ldots, 100))$ 

where d is computed as an aggregation



### Application II The Choquet integral in metric learning: reidentification

- Re-identification. Record linkage for databases, supervised approach
  - $\circ$  ML/Optimization for distance-based RL (A and B aligned).
    - ▷ Goal: as many correct reidentifications as possible:

for each record *i*, we need  $d(a_i, b_j) \ge d(a_i, b_i)$  for all *j* 



 $a_i = (a_{i1}, \ldots, a_{in})$  and  $b_i = (b_{i1}, \ldots, b_{in})$ 

- Re-identification. Record linking for databases. Supervised approach
  - ML/Optimization for distance-based approach. (A and B aligned)
    ⊳ Goal: as many correct reidentifications as possible. But, if error for a<sub>i</sub>: K<sub>i</sub> = 1 and d(a<sub>i</sub>, b<sub>j</sub>)+CK<sub>i</sub> ≥ d(a<sub>i</sub>, b<sub>i</sub>) for all j
    ⊳ or, expanding d, C<sub>p</sub>(diff<sub>1</sub>(a<sub>i1</sub>, b<sub>j1</sub>),..., diff<sub>n</sub>(a<sub>in</sub>, b<sub>jn</sub>)+CK<sub>i</sub> ≥ C<sub>p</sub>(diff<sub>1</sub>(a<sub>i1</sub>, b<sub>i1</sub>),..., diff<sub>n</sub>(a<sub>in</sub>, b<sub>in</sub>))
    ◦ Formalization:

$$\begin{array}{ll} \text{Minimize} & \sum_{i=1}^{N} K_i \\ \text{Subject to:} \mathbb{C}_p(diff_1(a_{i1}, b_{j1}), \dots, diff_n(a_{in}, b_{jn})) - \\ & - \mathbb{C}_p(diff_1(a_{i1}, b_{i1}), \dots, diff_n(a_{i1}, b_{i1})) + CK_i > 0 \\ & K_i \in \{0, 1\} \\ & \text{Additional constraints according to } \mathbb{C} \end{array}$$

- Re-identification. Record linking for databases. Supervised approach
  - $\circ$  ML/Optimization for distance-based approach. (A and B aligned)  $\circ$  Formalization for CI

 $\begin{array}{ll} \text{Minimize} & \sum_{i=1}^{N} K_i \\ \text{Subject to:} & CI_{\mu}(diff_1(a_{i1}, b_{j1}), \dots, diff_n(a_{in}, b_{jn})) - \\ & - CI_{\mu}(diff_1(a_{i1}, b_{i1}), \dots, diff_n(a_{i1}, b_{i1})) + CK_i > 0 \\ & K_i \in \{0, 1\} \\ & \text{Additional constraints for } \mu \end{array}$ 

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 $\begin{array}{ll} \text{Minimize} & \sum_{i=1}^{N} K_i \\ \text{Subject to:} & CI_{\mu}(diff_1(a_{i1}, b_{j1}), \dots, diff_n(a_{in}, b_{jn})) - \\ & - CI_{\mu}(diff_1(a_{i1}, b_{i1}), \dots, diff_n(a_{i1}, b_{i1})) + CK_i > 0 \\ & K_i \in \{0, 1\} \\ & \text{Additional constraints for } \mu \end{array}$ 

#### (but also WM, OWA, and Bilinear distance)

# Zooming out: trying to understand

Aggregation, distances, and independence

#### • Aggregation and distance.

Ο

- Arithmetic mean (AM): Euclidean distance
- Weighted mean (WM): Weighted euclidean
- Choquet integral (CI): Choquet integral-based distance
  - : Bilinear/Mahalanobis distance
- In a single picture: Mahalanobis and Choquet distance



#### Aggregation, distance and independence.

#### Only with Choquet integral and Mahalanobis distances

- ▷ Mahalanobis: covariance matrix
- Choquet integral: fuzzy measure
- In a single framework: Mahalanobis and Choquet *distance*



# Filling gaps:

#### Aggregation, distances, and independence

- Mahalanobis distance.
  - $\circ\,$  between  $\mathbf{x}\in\mathbb{R}^d$  and a vector  $\mathbf{m}\in\mathbb{R}^d$  with respect to the covariance matrix  $\Sigma$

$$(\mathbf{x} - \mathbf{m})\Sigma^{-1}(\mathbf{x} - \mathbf{m}))$$

- Choquet integral distance.
  - $\circ$  between  $\mathbf{x}\in \mathbb{R}^d$  and a vector  $\mathbf{m}\in \mathbb{R}^d$  with respect to a non-additive measure  $\mu$

$$CI_{\mu}((\mathbf{x}-\mathbf{m})\circ(\mathbf{x}-\mathbf{m})))$$

 $\mathbf{v} \circ \mathbf{w}$  is the Hadamard or Schur (elementwise) product of  $\mathbf{v}$  and  $\mathbf{w}$ (i.e.,  $(\mathbf{v} \circ \mathbf{w}) = (v_1 w_1 \dots v_n w_n)$ ).

#### • Choquet-Mahalanobis integral distance.

 $\circ \text{ between } \mathbf{x} \in \mathbb{R}^d \text{ and a vector } \mathbf{m} \in \mathbb{R}^d$  with respect to  $\mu$  and a positive-definite matrix Q

$$CMI(\mathbf{m}, \mu, \mathbf{Q}) = CI_{\mu}(\mathbf{v} \circ \mathbf{w})$$

where

- ▷  $\mathbf{L}\mathbf{L}^T = \mathbf{Q}$  is the Cholesky decomposition of the matrix  $\mathbf{Q}$ , ▷  $\mathbf{v} = (\mathbf{x} - \mathbf{m})^T \mathbf{L}$ ,
- $\triangleright w = \mathbf{L}^T(\mathbf{x} \mathbf{m})$ , and where
- $\triangleright \mathbf{v} \circ \mathbf{w}$  is the Hadamard (elementwise) product of  $\mathbf{v}$  and  $\mathbf{w}$ .

### **Choquet integral based distribution: generalized** *distance*

#### Well defined when $\Sigma$ is a covariance matrix.

 When Σ<sup>-1</sup> is a definite-positive matrix, the Cholesky descomposition is unique. This is the case when Σ is a covariance matrix valid for generating a probabilitydensity function.

#### **Proper generalization:**

- Generalization of both the Mahalanobis and the Choquet integral based distance.
  - The definition with  $\Sigma$  equal to the identity results into the Choquet integral of  $(x \bar{x}) \otimes (x \bar{x})$  with respect to  $\mu$ .
  - The definition with  $\mu$  corresponding to an additive probability  $\mu(A) = 1/|A|$  results into 1/n of the Mahalanobis distance with respect to  $\Sigma$ .

#### • Aggregation and distance.

Ο

- Arithmetic mean (AM): Euclidean distance
- Weighted mean (WM): Weighted euclidean
- Choquet integral (CI): Choquet integral-based distance
  - : Bilinear/Mahalanobis distance
- Choquet-Mahalanobis integral: CMI-distance



### A natural construction:

**Distributions** 

### Distributions

- E.g. in Classification data drawn from normal Gaussian distributions.
  - $\circ$  Parameters  $N(\mu,\Sigma)$  determined from real data or known
  - Set of k classes  $\Omega = \{\omega_1, \ldots, \omega_k\}$
  - $\circ$  covariance matrices  $\Sigma_i$
  - $\circ$  means  $\bar{x}_i$

class-conditional probability-density function Gaussian distribution

 $P(x|\omega_i) = \frac{1}{(2\pi)^{m/2} |\Sigma_i|^{1/2}} e^{-\frac{1}{2}(x-\bar{x}_i)^T \Sigma_i^{-1}(x-\bar{x}_i)}$ 



- Define distributions based on the Choquet integral. Why?
  - $\circ\,$  Non-additive measures on a set X permit us to represent interactions between objects in X !!
    - ... similar to covariances but different types of interactions !!

#### **Definition:**

- $Y = \{Y_1, \dots, Y_n\}$  random variables;  $\mu : 2^Y \to [0, 1]$  a non-additive measure and **m** a vector in  $\mathbb{R}^n$ .
- The exponential family of Choquet integral based class-conditional probability-density functions is defined by:

$$PC_{\mathbf{m},\mu}(\mathbf{x}) = \frac{1}{K} e^{-\frac{1}{2}CI_{\mu}((\mathbf{x}-\mathbf{m})\circ(\mathbf{x}-\mathbf{m}))}$$

where K is a constant that is defined so that the function is a probability, and where  $\mathbf{v} \circ \mathbf{w}$  denotes the Hadamard or Schur (elementwise) product of vectors  $\mathbf{v}$  and  $\mathbf{w}$  (i.e.,  $(\mathbf{v} \circ \mathbf{w}) = (v_1w_1 \dots v_nw_n)$ ).

#### Notation:

• We denote it by  $C(\mathbf{m}, \mu)$ .

### Distributions

#### • Shapes (level curves)



(a)  $\mu_A(\{x\}) = 0.1$  and  $\mu_A(\{y\}) = 0.1$ , (b)  $\mu_B(\{x\}) = 0.9$  and  $\mu_B(\{y\}) = 0.9$ , (c)  $\mu_C(\{x\}) = 0.2$  and  $\mu_C(\{y\}) = 0.8$ , and (d)  $\mu_D(\{x\}) = 0.4$  and  $\mu_D(\{y\}) = 0.9$ .

#### **Property:**

The family of distributions N(m, Σ) in ℝ<sup>n</sup> with a diagonal matrix Σ of rank n, and the family of distributions C(m, μ) with an additive measure μ with all μ({x<sub>i</sub>}) ≠ 0 are equivalent.
 (μ(X) is not necessarily here 1)

Follows from additivity in  $\mu = \text{probability} = \text{diagonal } \Sigma$
#### **Property:**

The family of distributions N(m, Σ) in ℝ<sup>n</sup> with a diagonal matrix Σ of rank n, and the family of distributions C(m, μ) with an additive measure μ with all μ({x<sub>i</sub>}) ≠ 0 are equivalent.
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Follows from additivity in  $\mu = \text{probability} = \text{diagonal } \Sigma$ 

#### **Corollary:**

• The distribution  $N(\mathbf{0}, \mathbb{I})$  corresponds to  $C(\mathbf{0}, \mu^1)$  where  $\mu^1$  is the additive measure defined as  $\mu^1(A) = |A|$  for all  $A \subseteq X$ .

#### **Properties:**

- In general, the two families of distributions  $N({\bf m}, {\bf \Sigma})$  and  $C({\bf m}, \mu)$  are different.
- $C(\mathbf{m}, \mu)$  always symmetric w.r.t.  $Y_1$  and  $Y_2$  axis.



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- Using the CMI distance, we consider both types of interactions
  Mahalanobis: Σ
  - Choquet (measure):  $\mu$

#### **Definition:**

- $Y = \{Y_1, \ldots, Y_n\}$  random variables,  $\mu : 2^Y \to [0, 1]$  a measure, **m** a vector in  $\mathbb{R}^n$ , and Q a positive-definite matrix.
- The exponential family of Choquet-Mahalanobis integral based classconditional probability-density functions is defined by:

$$PCM_{\mathbf{m},\mu,\mathbf{Q}}(x) = \frac{1}{K} e^{-\frac{1}{2}CI_{\mu}(\mathbf{v} \circ \mathbf{w})}$$

where K is a constant that is defined so that the function is a probability, where  $\mathbf{L}\mathbf{L}^T = \mathbf{Q}$  is the Cholesky decomposition of the matrix  $\mathbf{Q}$ ,  $\mathbf{v} = (\mathbf{x} - \mathbf{m})^T \mathbf{L}$ ,  $w = \mathbf{L}^T (\mathbf{x} - \mathbf{m})$ , and where  $\mathbf{v} \circ \mathbf{w}$  denotes the elementwise product of vectors  $\mathbf{v}$  and  $\mathbf{w}$ .

#### **Notation:**

• We denote it by  $CMI(\mathbf{m}, \mu, \mathbf{Q})$ .

#### **Property:**

- The distribution  $CMI(\mathbf{m},\mu,\mathbf{Q})$  generalizes the multivariate normal distributions and the Choquet integral based distribution. In addition
  - A  $CMI(\mathbf{m}, \mu, \mathbf{Q})$  with  $\mu = \mu^1$  corresponds to multivariate normal distributions,
  - $\circ \ {\rm A} \ CMI({\bf m},\mu,{\bf Q}) \ {\rm with} \ Q = \mathbb{I} \ {\rm corresponds} \ {\rm to} \ {\rm a} \ CI({\bf m},\mu).$

#### **Graphically:**

• Choquet integral (CI distribution), Mahalobis distance (multivariate normal distribution), generalization (CMI distribution)



1st Example: Interactions only expressed in terms of a measure.

- No correlation exists between the variables.
- CMI with  $\sigma_1 = 1$ ,  $\sigma_2 = 1$ ,  $\rho_{12} = 0.0$ ,  $\mu_x = 0.01$ ,  $\mu_y = 0.01$ .



**2nd Example:** Interactions only in terms of a covariance matrix.

• CMI with  $\sigma_1 = 1$ ,  $\sigma_2 = 1$ ,  $\rho_{12} = 0.9$ ,  $\mu_x = 0.10$ ,  $\mu_y = 0.90$ .



**3rd Example:** Interactions both: covariance matrix and measure.

• CMI with  $\sigma_1 = 1$ ,  $\sigma_2 = 1$ ,  $\rho_{12} = 0.9$ ,  $\mu_x = 0.01$ ,  $\mu_y = 0.01$ .



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#### **Example:**

- Non-additive  $\mu$ :  $CMI(\mathbf{m}, \mu, \mathbf{Q})$  not repr. spherical/elliptical
- No *CMI* for the following spherical distribution: Spherical distribution with density

$$f(r) = (1/K)e^{-\left(\frac{r-r_0}{\sigma}\right)^2},$$

where  $r_0$  is a radius over which the density is maximum,  $\sigma$  is a variance, and K is the normalization constant.

# **Summary**

# Summary

#### Summary:

- Choquet integral and non-additive measures for decision and reidentification
- Definition of distances based on the Choquet integral
- Comparison with the Mahalanobis distance
- Construction of distributions
- Relationship with multivariate normal and spherical distributions

# Thank you